A Simple Note on Informational Cascades^{*}

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Abstract

Seminal models of herd behaviour and informational cascades point out existence of negative information externalities, and propose to 'destroy' information in order to achieve social improvements. Although in the last years many features of herd behaviour and informational cascades have been studied, this particular aspect has never been extensively analysed. In this article we try to fill this gap, investigating both theoretically and experimentally whether and to which extent destroying information can improve welfare. Our empirical results show that this decisional mechanism actually leads to a behaviour pattern more consistent with the theory that in turn produces the predicted efficiency gain.

Keywords: Informational cascades; Information Externality; Individual Decision Making; Experiment *JEL Classification:* C91; D62

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1 Introduction

Part of social learning is related to an apparently naive behaviour known as *herd behaviour* (Banerjee, 1992)¹ or *informational cascades* (Bikhchandani, Hirshleifer, and Welch, 1992 – BHW, henceforth). A peculiarity of these models, however, is that they view agents' imitative behaviour as perfectly rational, even though characterized by imperfect information.

This behaviour takes place when agents can augment their information set by looking at other agents' behaviour. Although rational, it could cause information externalities that result in an aggregate welfare loss (Becker, 1991). In this situation, the individual rational behaviour may well result in a non-optimal strategy from an aggregate point of view. Looking at the real world, we have abundant empirical evidence for informational cascades. Actually, one of the most attractive features of these kinds of models concerns their direct application to a range of every-day situations. Just to cite an example, we can refer to bubbles in financial markets (Plott, 2002; Hey and Morone, 2004; Morone, 2008). The idea underlying these models is simple. Consider the case in which somebody has to choose between two unknown restaurants and he/she has no relevant information about them. However, he/she can infer that the most crowded is the best one and will choose to join the queue. This behaviour is rational, but the possibility that first customers have no pregnant information as well is crucial. BHW point out that the conformity of followers in a cascade contains no informational value (p. 998-999), and this argument has been demonstrated by some empirical evidence (Anderson and Holt, 1997; Allsopp and Hey, 2000). On the other hand, also the impact of signal accuracy

¹ A generalization of Banerjee's model was proposed in Morone and Samanidou (2008).

on cascade efficiency has been demonstrated by some simulations (e.g., Pastine and Pastine, 2005). From a theoretical perspective, Sgroi (2002) has evaluated the strategy of forcing a subset of individuals to make their decisions without observing others' actions. The aim of this paper lies in the path opened by Sgroi (2002). It is investigating (experimentally) the possibility to mitigate informational cascades' negative effects forcing the first *k* subjects in a queue to play only according to their private information. For this purpose, we analyse a sequential model departing from BHW's model in some relevant parts and then we experimentally investigate if *"society may actually be better off by constraining some of the people to use only their own information*" (Banerjee, 1992; p. 798).

The paper is structured as follows. Section 2 is devoted to the new specification of the standard model. The experimental design and results are introduced, respectively, in Sections 3 and 4. Section 5 concludes.

2 Theory

In addition to the seminal papers on herd behaviour and informational cascades (Banerjee, 1992; Bikhchandani et al., 1992), even more recently a paper on word-ofmouth learning (Banerjee and Fudenberg, 2004) notes that the inefficient herding of the standard models does not occur if some agents are forced to use their own private information. Sgroi (2002) considers gains in using a group of "guinea pigs" (agents that move simultaneously to subject 1) from the perspective of a social planner. Sgroi, as a first step, develops individual payoff functions, and then a total consumer welfare function, $\Omega(M N, p)^2$. For example, he shows that $\Omega(1, N, p) > \Omega(0, N, p)$, but that $\Omega(N-1, N, p) < \Omega(0, N, p)$, demonstrating that the optimal level of M will lie strictly between 0 and N-1. Moreover, his results on the total consumer welfare function Ω can be used to calculate the unique optimal number of agents that play simultaneously on the base of their own private information, and the percentage improvement over the standard situation, i.e., where all agents can see others' actions.

Our approach has been different: we provide a complete generalization of BHW's model, allowing for a generic subset k subjects that make decisions without occasion of social learning, and then we calculate a new set of herd probabilities. In order to make empirical comparison with the standard model, we do consider the percentage of winning as a useful proxy for individual utility.

We have a population of $I = \{1, ..., N\}$ individuals. Each individual $i \in N$ has to decide whether to adopt a specific behaviour, for example, whether to adopt a new technology or not. All individuals make their choices in a sequential and exogenously determined order. The gain of adopting, V, is the same for all $i \in N$ and is either zero or one. These two events have the same *ex-ante* probability to occur.

However, each individual *i* privately observes a conditionally i.i.d. signal about *V*. This signal *s* is either 0 or 1: 1 is observed with probability $p > \frac{1}{2}$ if the true value is 1, and with probability *1-p* otherwise.

² In brackets, M represents number of guinea pigs, N number of individuals in the population, p the quality of signals.

Under our specification, the first k (< N) individuals in the queue are not allowed to observe the decisions already taken, whereas the entire history of decisions is commonly known to the last N-k individuals. We can think of this game as of N-stage game where the first k individuals play simultaneously and the remaining N-k sequentially. As the first k individuals can observe only their own signal, rationality requires them to follow their private information: they should take on the new behaviour if the signal is 1, and reject it otherwise. In contrast, the remaining N-k individuals should base their decision on both their own signal and all past decisions, thereby choosing the most frequently observed action³. In case of indifference, we assume that individual i with i = k+1, ..., N follows the tie-breaking rule of adopting or rejecting with equal probability.

In our specification it is as if individual *i*, with i = k+1, ..., N has an advantage of additional signals. In this manner, we expect our specification to lead to a more socially efficient final outcome, as the society has a mechanism that allows aggregating the information in a later stage and in a more correct way.

In their model (T1), where all decision makers are allowed to observe their predecessors' action, BHW derive the unconditional *ex ante* probability of a cascade and the *ex ante* probability of no cascade after an even number of individuals *n*.

They also derive the probabilities of ending up in a correct cascade and ending up in a wrong one. We derive the same probabilities after having taken in consideration the fact that the first k players act only based on their own signal s

³ More precisely, Anderson and Holt (1997) show that the optimal strategy in a Bayesian sense whenever the two events are equally probable and signals identically distributed corresponds to the very simple strategy of doing the count of the previous decisions, one's own signal included.

(T2). We show our main results in the Appendix A. At this point, however, it may be more illustrative to compare probabilities of ending up in a correct cascade (to a some extent, it can be considered as an index of efficiency) under the two specifications for some different parameter values (k, number of players observing only their own signal; p, probability of signal correctness; n, number of players having already taken their decision). Figures are shown in Table 1.

	<i>p</i> = 0.	75 ; <i>n</i> = 10	00		(k = 6); p =	0.75		(k = 56); n = 100					
	(k = 10)	(<i>k</i> = 56)	(<i>k</i> = 98)	<i>n</i> = 10	<i>n</i> = 100	<i>n</i> = 1000		<i>p</i> = .55	<i>p</i> = .85	<i>p</i> = .99			
т1	0.8077	0.8077	0.8077	0.8075	5 0.8077	0.8077	_	0.5664	0.9011	0.9949			
т2	0.969	0.9999	0.9999	0.9333	3 0.937	0.937		0.7778	0.9999	1			

Table 1 Probability of a correct cascade: Comparative static analysis and comparison between models⁴

At a first glance, it is evident that probability of ending up in a correct cascade is higher in T2. Entering into details, we can point out that as k increases (left panel), probability of a correct cascade becomes not statistically different from 1, whereas under the standard model the probability is quite high, but never reaches this level. Other results are more obvious, in the sense that probability of a correct cascade is monotonically increasing in the number of subjects that have already made a decision (n, middle panel) and in the signal correctness (p, right panel) under both the two specifications, but nevertheless always higher under ours.

Probably, it may be interesting to combine results regarding the effect of changes in signal correctness and number of subjects that act with no clue regarding

⁴ In the first row are reported different parameter values at which probabilities are computed and in boldface, parameters held constant for each comparative static exercise. Different values of k are in parentheses since relevant only to our model.

previous decisions. We perform these comparative static exercises varying simultaneously p and k, while keeping N - the number of individuals in the population – constant. Consequently, we have the opportunity to note that for each value of k there is a probability p^* , under which the difference between the two specifications is maximised. The converse is also true.

3 Experimental design

In order to test empirically whether the new specification of the model allows achieving a social improvement, we ran a computerized experiment at the laboratory of ESSE at the University of Bari. The experiment was composed by two treatments: the control treatment (T1) was set in accordance with the original model, whereas in the second treatment (T2) we test the new specification, with the first k = 4 subjects forced to play basing their decision exclusively on their private information⁵. Both treatments was composed by two sessions.

The experiment was programmed using the Z-tree software (Fischbacher, 2007). Each treatment lasting for about an hour was made up of 22 periods, of which 2 were trial ones. The trial periods were necessary for subjects to become familiar with the experiment, providing them also the opportunity to ask questions about the instructions (Appendix B). The final payment was made only for the 20 real periods and paid at the end of each treatment.

⁵ As we noted above, there is an optimal value for k^* for each parameter combination. We determined the optimal k^* with a Monte Carlo simulation. This simulation provided the winning percentages for each position in the queue, provided that we consider the individual winning percentage as a proxy for individual utility. The simulation consisted of 10 millions iterations for each different value of k, setting N and p at 10 and .75, respectively. We get a measure for social welfare summing up individual winning percentages over the entire population, and we picked the case in which this indicator was at its maximum.

We had N = 10 subjects for each session (four in total) sitting next to a PC terminal connected by a net. The subjects could not see each other or communicate. All of them were students of Economics not familiar with previous similar experiments. In the experiment, subjects acted as entrepreneurs and their task was to decide whether to invest in a new product or not. The order in which they chose sequentially was randomly determined period by period⁶. However, subjects did not know whether this product would be profitable or not once on the market. Whenever they made the right decision, they gained $\notin 0.5$, and zero otherwise⁷. For each period the programme established the true value of V but did not reveal it to subjects. Each of them, however, received a free-of-charge signal s about V(a sort of f)a result of a market survey). These signals took either the value 1 or the value 0 and the signal correctness (p = .75) was common knowledge. The screen displayed these details: one's own turn to play; the position in the queue; where allowed, the decision made by predecessors; and one's own signal. At the end of each period, subjects were informed about the right option and their payoff. When all periods were played, subjects were paid and free to leave the laboratory. Average payoff was €7.125.

4 Results

We start this section showing some data at individual level (raw data provided in the Appendix C). In each position, taking into account predecessors' decisions and the

⁶ They were informed about their turn via a message on their PC screen.

⁷ More precisely, if the product was successful (V = 1), they would gain $\notin 0.5$ in case of investment, and zero otherwise. If the product was not successful (V = 0), they would gain $\notin 0.5$ in case of no investment (the right decision in this scenario), and zero otherwise.

signal realization, we determine which action should be chosen according to the theory. Consequently, we categorize as rational behaviours in accordance with it (in our simple set-up, the optimal strategy in a Bayesian sense corresponds to the count, as explained in footnote 3) and, in case of indifference, whenever subjects adopted the tie-breaking rule, regardless of the fact that it produces a cascade or not. In particular, at the individual level, in the table we report also as cascade behaviour all the cases where "an imbalance of previous inferred signals causes a person's optimal decision to be inconsistent with his or her private signal" (Anderson and Holt, 1997; p. 851), that is, all the cases in which individuals follow the queue, disregarding their signal. As regards behaviours categorized as irrational, namely, inconsistent with the theory, we discriminate cases in which it can be rationalized somehow – following her own signal – from cases where it cannot be explained whatsoever. Results are reported in Table 2.

			Rational Beha	viour	Irrational	behaviour
	-		Occurrence of	rational cascades	not rationalized	signal-keeping
			correct	wrong		
on 1	т1	146	4	16	33	21
Session 1	т2	171	12	8	24	5
Session 2	т1	162	16	11	15	23
Sessi	т2	177	14	1	7	16
ment	т1	308	20	27	48	44
Experiment	т2	348	26	9	31	21

Table 2 Summary of behaviours observed in the experiment (individual level)

From table 2 it can be clearly noted that the new specification may be effective in driving a more consistent behaviour (308 out of 400 observations - 77% - in T1 vs. 348 out of 400 observations - 87% - in T2). However, though in absolute terms occurrence of irrational behaviours in T2 is lower (92 cases in T1 vs. 52 cases in T2), percentage of behaviour that cannot be explained in any case is higher than in T1 (48 out of 92 cases in T1 - 52.2% - vs. 31 out of 52 cases in T2 - 59.6%). Moreover, it is interesting to note that not only occurrence of cascade behaviour under the two treatments is significantly different (47 cases in T1 vs. 35 case in T2), but also percentage of a correct cascade is almost two times higher in T2. From the table it is also simple to note that results are very similar across sessions.

Finally, we observe that cascade behaviour is rather fragile (individuals do not choose to conform to the mass, still when their all predecessors made the same choice)⁸, and that often they also choose to play against their own signal, especially when they are the first in the queue⁹.

At this point, in order to test our hypothesis, namely that under the new specification of the model the outcome is socially more efficient, we compare the average earnings under the two treatments.

Particularly interesting is the comparison between the theoretical earnings, as it would have been if all individuals behaved according to the theory, given the actual signal realization during the experiment, and the experimental earnings, the

⁸ As an extreme case of this kind of behaviour, we can cite as example the behaviour of subject in position 10 (period 10) under T1 in Session 1 that decided to break the cascade even if all the players before her made the same decision.

⁹ In particular, we observe that first individual in the queue played against their own signal in 8 cases out of 20 periods in T1 in Session 1 and in 2 cases in Session 2. Considering that in T2 the first four players made decisions without observing predecessors' signal, we observe this kind of behaviour in 17 instances in Session 1 and in 5 cases in Session 2.

actual payoffs obtained by participants during the experiment. Results are reported in Table 3^{10} .

First, we note a statistically significant difference between the two treatments (Wilcoxon rank-sum test for theoretical earnings: -3.863, *p*-value = .0001; for experimental earnings: -3.780, *p*-value = .0002). Interestingly, each experimental treatment is not statistically different from its theoretical counterpart (Wilcoxon rank-sum test for BHW and T1: 1.625, *p*-value = .1043; for our specification and T2: 1.359, *p*-value = .1743). Second, for each position in the queue, we observe always higher average earnings under T2¹¹.

Position in the	Theoretic	al earnings		Experime	ental earnings	
queue	т1	т2	Т	1		т2
			S1	S2	S1	S2
1	0.225	0.4	0.225	0.4	0.356	0.394
2	0.325	-	0.275	0.375	-	-
3	0.3125	-	0.325	0.3	-	-
4	0.33125	-	0.275	0.3	-	-
5	0.34375	0.4	0.275	0.4	0.425	0.4
6	0.34375	0.4125	0.25	0.325	0.375	0.35
7	0.34375	0.4125	0.3	0.35	0.375	0.45
8	0.34375	0.40625	0.3	0.35	0.4	0.4
9	0.34375	0.425	0.375	0.35	0.375	0.45
10	0.34375	0.425	0.375	0.275	0.4	0.45
TOTAL	3.25625	4.08125	2.975	3.425	3.775	4.076

¹⁰ It is important to note that the probability to obtain a correct signal was identical across the treatment (fixed at 0.75 over all the treatments), but that the actual realization of signal was different. Consequently, since chance rather than behaviour could explain the difference between the two treatments, we standardise our results dividing the average earnings in Table 3 by the frequency of correct signals, position by position. Results in Appendix D.

¹¹ For this result, we averaged earnings between different sessions under the two treatments.

¹² Since the first four positions in T2 can be considered as theoretically indistinct, we decided to provide in the table an average of the relative earnings.

In order to give an additional insight into the social efficiency gain, we can consider the percentage of winning as a useful proxy for individual utility, and then compare this index under the two treatments. In figure 1 we report the index for each different position held in the decisional queue.

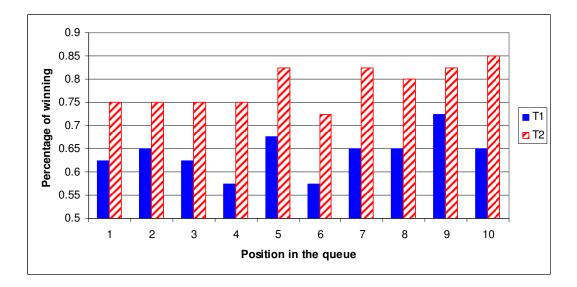


Figure 1 Percentages of winning¹³

It is clear that under the second treatment the percentages of winning are always higher than percentages under the first treatment. Looking at the graph, we may state that the new decisional mechanism is preferable from a social and even individual point of view, i.e. in T2 the winning percentages is always larger then in T1.

¹³ As explained in footnote 12, we averaged percentages the first four positions in T2 given that they can be considered theoretically indistinct.

4.1 Econometric analysis

Finally, we estimate a very simple learning model. Specifically, we constructed a model that links decisions in the experiment to a set of determinants, as follows. First, the presence of learning is investigated by the use of the variable *Time* (the period number) and the variable *Time*², to test for concavity of learning. Moreover, in order to gain further insight, we test for the presence of directional learning (or Cournot behaviour; Selten and Buchta, 1998) by using the *Correctwon* variable (a dummy variable equal to 1 if in the most recent period the subject made the theoretically correct decision and won) and *Correctlost* variable (a dummy variable equal to 1 if in the subject made the theoretically correct decision and won) and *Correct,* is a dummy variable equal to 1 whenever subjects make the decision consistent with the theory, 0 otherwise. Consequently, we run a probit estimation procedure. Results are reported in Table 4.

Dep. Variable: <i>Correct</i>	Marginal	Std. Error	p-value
	Effect		
Time	.01884	.01049	0.073
<i>Time</i> ²	00069	.00049	0.158
Correctwon	05218	.03423	0.118
Correctlost	13801	.06793	0.026
т2	.07620	.02924	0.009
Log likelihood	-409.62579		
Pseudo R ²	0.0311		
NOBs	800		

 Table 4 Maximum Likelihood probit estimation¹⁴

¹⁴ Note that the reported significance levels assume independent observations, though this is unlikely to be the case.

We note a moderate trend in observing a more consistent decisions over time (in fact, *Time* is positive); but concavity for learning is not statistically significant. We can affirm that directional learning is not a firm determinant of learning in our experiment. Indeed, even if *Correctlost* is significant and present correct direction, *Correctwon* is neither significant nor presents the theoretically correct direction (positive sign) in our analysis. In fact, we expect the probability of making the correct decision to increase if in the previous period subject made correct decision and won.

Interestingly, on the other hand, the dummy variable for the treatment T2: the decisional mechanism implemented in this treatment is actually effective in increasing probability of making the correct decision by 7,6%¹⁵.

5 Conclusions

Negative informational externality produced by phenomenon of informational cascade has drawn quite a lot of attention in economic literature. Hence, finding mechanisms useful in eliminating or at least minimising this externality are of quite an interest. The paradox whereby burning a piece of information in a first stage of the sequential decisional process could turn to be a social improvement in a later stage was indeed worth investigating. This was our task. Our empirical results show that this decisional mechanism actually leads to behaviour more consistent with the theory that in turn produces a social improvement. If supported by further analyses

¹⁵ However, it could be of interest to test if the dummy variable T2 has a different impact, depending on the different positions taken into account. In order to test for the presence of a possible *structural break* at stage 4, we perform a Chow test. Results are given in Appendix E.

 aimed, for example, to design an implementable self-enforcing mechanism - our result may open new challenging scenarios once applied to reality.

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Appendices

A - In order to get probabilities in the same fashion as BHW obtained, we derived them varying each time the value of k – number of individuals acting with no clue regarding previous actions. Then, having been noted some regularities, we generalized the model, whatever the value for k.

The probability of NO-cascade after n = 2m individuals is simply the probability of observing the same occurrence for the two signals, i.e., s = 0 and s = 1. Consequently, in our specification, the probability of NO-cascade is shown in 1.a, whenever k is an even number:

$$\frac{k!}{\left(\frac{k}{2}\right)!\left(\frac{k}{2}\right)!}p^m(1-p)^m$$
 1.a

whereas, whenever k is an odd number:

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$$\frac{1}{2} \left[\frac{(k+1)!}{\left(\frac{k+1}{2}\right)!} \right] p^m (1-p)^m.$$
 1.b

However, it is of greater importance to consider the probability of ending up in a CORRECT-cascade. The probability to observe a correct cascade after 2 individuals is simply the probability that both of them receive a correct signal, i.e., p^2 . The probability to observe a correct cascade after 4 individuals is simply the sum of the probabilities of the two independents events: the occurrence of a correct

cascade after 2 individuals OR the occurrence of no cascade after 2 individuals AND the occurrence to observe a correct cascade for the third and fourth players.

In general formulas, after n = 2m individuals, whenever k is an even number:

$$\sum_{j=\frac{k}{2}+1}^{k} \binom{k}{j} p^{j} (1-p)^{(k-j)} + \frac{k!}{\left(\frac{k}{2}\right)!} \binom{k}{2!} p^{\frac{k}{2}} (1-p)^{\frac{k}{2}} \frac{p(p+1)}{2} \left\lfloor \frac{1-(p-p^{2})^{m-\frac{k}{2}}}{1-(p-p^{2})} \right\rfloor$$
 2.a

and whenever k is an odd number:

$$\left(\frac{1}{4}\binom{k+1}{\frac{k+1}{2}+1}p^{j}(1-p)^{(k+1-j)} + \sum_{j=\frac{k+1}{2}+2}^{k+1}\binom{k+1}{j}p^{j}(1-p)^{(k+1-j)}\right) + \frac{1}{2}\left[\frac{(k+1)!}{\left(\frac{k+1}{2}\right)!\binom{k+1}{2}!}\right]p^{\frac{k+1}{2}}(1-p)^{\frac{k+1}{2}}\frac{p(p+1)}{2}\left[\frac{1-(p-p^{2})^{m-\frac{k+1}{2}}}{1-(p-p^{2})}\right] 2.b$$

Finally, a similar reasoning can be applied to calculate the probability of ending up in a WRONGcascade after n = 2m individuals. In formulas, whenever k is an even number:

$$\sum_{j=\frac{k}{2}+1}^{k} \binom{k}{j} (1-p)^{j} p^{(k-j)} + \frac{k!}{\left(\frac{k}{2}\right)!} \frac{p^{\frac{k}{2}}(1-p)^{\frac{k}{2}} \frac{(p-2)(p-1)}{2}}{2} \left[\frac{1-(p-p^{2})^{m-\frac{k}{2}}}{1-(p-p^{2})} \right]$$
 3.a

and whenever k is an odd number:

$$\left(\frac{1}{4}\binom{k+1}{\frac{k+1}{2}+1}(1-p)^{j}p^{(k+1-j)} + \sum_{j=\frac{k+1}{2}+1}^{k+1}\binom{k+1}{j}(1-p)^{j}p\right) + \frac{1}{2}\left[\frac{(k+1)!}{(\frac{k+1}{2})!\binom{k+1}{2}!}\right]p^{\frac{k+1}{2}}(1-p)^{\frac{k+1}{2}}\frac{(p-2)(p-1)!}{2}\left[\frac{1-(p-p^{2})^{m-\frac{k+1}{2}}}{1-(p-p^{2})!}\right] 3.b$$

To make an example, we provide the calculations for determination of probability of NO-cascade, of ending up in a CORRECT-cascade or in a WRONG-cascade for the specific value of parameters chosen for the experiment, i.e., k = 4, m = 5, and p = 0.75.

Applying formula 1.a:

$$\frac{4!}{2!2!}(0.75)^5(0.25)^5 = 0.001390$$

Applying formula 2.a:

$$\sum_{j=3}^{4} \binom{4}{j} 0.75^{j} (0.25)^{(4-j)} + \frac{4!}{2!2!} (0.75)^{2} (0.25)^{2} \frac{0.75(1.75)}{2} \left[\frac{1 - (0.75 - 0.5625)^{3}}{1 - (0.75 - 0.56.25)} \right] = 0.907531$$

Finally, considering formula 3.a, we get probability of ending up in a WRONG-cascade:

$$\sum_{j=3}^{4} \binom{4}{j} (0.75)^{j} (0.25)^{(4-j)} + \frac{4!}{2!2!} p^{2} (0.25)^{2} \frac{(-1.25)(-0.25)}{2} \left[\frac{1 - (0.75 - 0.5625)^{3}}{1 - (0.75 - 0.5625)} \right] = 0.091079$$

B - Instructions (original provided in Italian)

Welcome! This experiment is designed to study how people make decisions. The experiment is very simple, and you will have the possibility of earning money, which will be paid to you in cash at the end of the experiment.

This amount will depend, on the one hand, on your decisions and, on the other hand, on luck.

You will play as an entrepreneur and your task will be to decide to develop a new product or not.

Two scenarios will have the same probability to occur: or all goods will be sold or not a single one.

You will repeat your task 20 times. In each period, the computer will choose the scenario. The scenario will be the same for all the participant, but different in each period.

Whenever you take the right decision, you will earn 0.5€, nothing otherwise, as shown in the table:

	Decision: to invest	Decision: not to invest
All goods sold	0.5€	0
No good sold	0	0.5€

It is important to know that you make your decision in sequence and the order is randomized in each period.

However, you will be provided with two different kinds of information before making your decision.

First, you will receive results from a market survey reliable at 75%. In particular, during the experiment you will be provided with a signal according to the result of the survey. As shown in the table, to each signal is connected a different likelihood of the two scenarios:

	Signal = 1	Signal = 0
All goods sold	75%	25%
No good sold	25%	75%

Second, you will be informed about decisions already made by all entrepreneurs before you.

You will not be required to pay for these pieces of information. These will appear automatically on your PC screen when it is your turn to play.

It is important to note that the first four players will not receive this second kind of information. On the contrary, from the fifth player onwards, players will receive all relevant information regarding previous decisions.

Whenever you make your decision, you have to press the OK button to confirm your choice. As soon an all players have made their decision, on your PC screen you will be informed about the right choice to take in that period and your relative payoff.

You will play for 20 periods, in addition to two trial periods at the beginning of the experiment.

At the end, you will be paid (except for the payoff earned during the trial periods) and you will be free to leave the laboratory.

The rules are very simple. However, please do not communicate with other participants during the experiment. You are free to put questions to experimenters at any time during trial periods raising your hands.

Good luck!

C – *Raw data*

Treatment 1 Session 1

Period	Subject	Signal	Action												
1	6	1	1	6	6	1	1	11	6	1	1	16	1	1	1
1	2	1	1	6	7	1	1	11	4	0	0	16	5	0	0
1	4	1	0	6	3	1	1	11	7	1	1	16	6	0	0
1	7	1	1	6	9	1	1	11	10	0	0	16	10	0	0
1	3	1	1	6	1	0	0	11	9	0	1	16	8	0	0
1	1	1	1	6	8	1	1	11	8	0	0	16	9	0	0
1	9	1	1	6	10	1	1	11	5	0	0	16	7	1	0
1	8	1	1	6	4	1	1	11	3	0	1	16	2	0	0
1	5	1	0	6	2	1	0	11	1	0	0	16	3	1	1
1	10	1	1	6	5	1	1	11	2	0	1	16	4	0	0
2	3	0	1	7	7	0	0	12	2	0	0	17	3	1	1
2	9	0	0	7	1	1	0	12	1	0	0	17	2	0	1
2	5	1	1	7	2	0	0	12	9	1	1	17	8	1	1
2	4	0	1	7	6	0	1	12	4	0	1	17	5	0	0
2	10	1	1	7	4	Õ	0	12	10	1	1	17	1	0	0 0
2	6	1	1	7	3	Õ	0	12	8	1	1	17	6	Ő	0
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4	10	0	0	9	7	0	Ő	14	6	0	1	19	10	Ő	Ő
5	8	0	1	10	2	Ő	0	15	5	0	1	20	8	0	0
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Ũ				10					8			20	10	1	1
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Treatment 1 Session 2

Period	Subject	Signal	Action	Period	Subject	Signal	Action	Period	Subject	Signal	Action	Period	Subject	Signal	Action
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	2	1	0	6	1	0	0	11	2	1	1	16	1	1	1
1	6 10	0	0 1	6	2 10	1	0	11 11	5 8	1	1	16	7 9	0 1	0
1	5	1	1	6 6	4	1	1 0	11	8	0	1	16 16	9 10	1	1
1	э 4	1	1			-				1	1			1	0
1	4 7	1	1	6	6 7	1	0	11	10 7		0	16	2	1	1
1	3	1	1	6	7 5	1	1 0	11 11		0	0	16 16	8 6	1 0	0
1				6					3		1				1
1	8	1	1	6	3	0	0	11	6	0		16	3	1	1
2	2	0	0	7	7	1	1	12	5	•	1	17	2	1	1
2 2	8 10	0 0	0 0	7 7	5 6	1 0	1 0	12 12	8 7	1	1	17	3 8	1 0	0
										•	-	17			0
2	5	0	0	7	4	0	0	12	1	1	1	17	9	1	1
2	6	1	0	7	10	1	1	12	2	0	1	17	1	1	1
2	4	1	1	7	3	0	0	12	6	1	1	17	5	1	1
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3	10	1	1	8	6	0	0	13	1	1	1	18	9	1	0
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4	2	0	0	9	1	0	0	14	6	0	0	19	2	1	0
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4	10	0	0	9	6	0	0	14	3	0	0	19	10	1	1
4	4	0	0	9	5	0	0	14	5	0	0	19	7	0	0
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Treatment 2 Session 1

Period	Subject	Signal	Action	Period	Subject	Signal	Action	Period	Subject	Signal	Action	Period	Subject	Signal	Action
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1	6	0	0	6	3	0	0	11	3	0	0	16	2	1	1
1	4	0	0	6	1	0	0	11	8	0	1	16	6	1	1
1	2	0	1	6	4	0	1	11	4	0	1	16	8	1	1
1	5	0	0	6	6	0	0	11	6	1	1	16	4	1	0
1	7	0	0	6	9	0	0	11	7	1	1	16	7	1	1
1	9	0	0	6	8	0	0	11	1	0	0	16	3	1	1
1	1	0	0	6	5	0	0	11	5	0	0	16	5	1	1
1	8	0	0	6	10	0	0	11	9	1	1	16	9	0	1
2	2	1	1	7	4	1	1	12	6	0	0	17	5	0	0
2	6	0	0	7	5	1	1	12	2	0	0	17	9	1	0
2	5	1	0	7	7	1	1	12	1	0	0	17	10	0	0
2	1	0	0	7	1	1	1	12	5	0	0	17	2	1	0
2	10	0	1	7	8	0	1	12	8	0	0	17	6	1	1
2	9	0	1	7	2	1	1	12	9	0	0	17	1	1	1
2	8	1	1	7	6	1	1	12	10	Õ	õ	17	3	1	0
2	7	0	1	7	10	1	1	12	7	Õ	õ	17	4	1	1
2	3	Õ	1	7	9	1	1	12	3	Õ	õ	17	8	0	0
2	4	1	1	7	3	1	1	12	4	Ő	1	17	7	0	0 0
3	7	0	1	8	2	1	0	13	8	0	0	18	3	0	0
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	1			8	8 7	1	1	13	4		1	18			-
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3	5	0	0	8	4	1	1	13	10	0	0	18	1	0	0
3	6	0	1	8	10	1	1	13	9	0	0	18	8	1	0
3	4	0	0	8	5	0	0	13	5	0	0	18	2	0	1
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4	8	1	0	9	1	0	0	14	1	0	1	19	1	1	1
5	8	1	1	10	10	0	1	15	7	1	1	20	6	0	0
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5	2	1	1	10	5	0	1	15	1	1	1	20	10	1	0
5	2	I	I	10	э	U	I	15	I	I	I	20	10	I	U

Treatment 2 Session 2

1 2 1 1 11 1	Period	Subject	Signal	Action												
1 7 1 1 1 11 5 1 1 16 6 1 1 1 4 1 1 6 2 1 1 14 1 16 5 1 1 1 4 1 1 6 2 1 1 11 6 1 1 16 3 0 0 1 4 1 1 6 1 1 1 11 10 1 16 16 1 1 11 17 16 1 16 1 16 1 1 1 11 11 17 10 1 1 1 1 1 11 17 10 1 1 1 1 1 1 1 11 17 1 1 1 1 1 11 17 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	1														-	1
1 9 1	1										-					0
1 4 1 1 6 2 1 1 11 6 1 1 16 3 0 0 1 8 0 0 6 10 1 1 11 3 1 1 16 7 0 0 1 5 0 0 6 7 1 1 11 3 1 1 16 8 0 0 1 10 1 1 6 6 1 1 11 17 0 1 16 8 0 0 2 7 0 0 7 10 1 1 12 10 1 17 10 1 1 2 1 1 7 2 1 1 17 2 1 1 1 1 1 17 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 <t< td=""><td>1</td><td></td><td></td><td></td><td></td><td></td><td>-</td><td></td><td></td><td></td><td>-</td><td></td><td></td><td></td><td>•</td><td>1</td></t<>	1						-				-				•	1
1 1	1			1			-								•	1
1 8 0 0 6 10 1 1 11 2 1 1 16 7 0 0 1 10 1 1 6 3 1 11 2 1 0 166 8 0 0 2 4 0 0 7 10 1 11 7 0 1 16 8 0 0 2 4 0 0 7 10 1 11 7 1 1 17 10 1 11 17 1 1 11 12 1 1 17 1 1 11 11 17 1 1 11 12 2 1 1 17 1 1 11 12 1 1 17 1 1 11 12 1 1 17 3 1 1 11 17 3 1 1 11 11 11 11 11 11 11 11 11	1			1			-				-					
1 5 0 6 7 1 11 11 2 1 1 16 1 1 0 16 1 1 0 1 6 1 1 11 11 7 0 1 16 2 1 0 2 4 0 0 7 6 0 0 12 2 1 1 17 5 1 1 2 7 1 1 7 6 0 0 12 2 1 1 17 5 1 1 2 1 1 7 6 0 0 12 2 1 1 17 1 1 1 2 1 1 7 4 0 0 12 3 1 1 11 12 1 1 1 17 1 1 1 2 3 0 0 7 7 1 1 1 1 1 1 <	1							1				-				-
1 10 1 1 6 3 1 1 11 7 0 16 8 0 0 2 4 0 0 7 10 1 12 10 1 1 17 5 1 1 2 7 1 1 7 2 1 1 12 2 1 1 17 5 1 1 2 2 0 0 7 2 1 1 17 5 1 1 1 1 17 5 1 1 1 12 8 1 1 12 9 0 1 177 7 1 1 2 9 0 0 7 3 1 1 12 3 1 1 177 3 1 1 2 6 0 0 7 7 1 1 12 3 1 1 1 1 1 1 1 1 1 <	1			-			-	1							-	-
1 6 1 1 11 7 0 1 16 2 1 0 2 7 1 1 7 6 0 12 2 1 1 17 5 1 1 2 7 1 1 7 6 0 0 12 2 1 1 17 5 1 1 2 1 1 1 7 4 0 0 12 8 1 1 17 9 1 1 2 8 1 1 7 1 1 12 7 1 1 17 8 0 0 2 9 0 0 7 3 1 1 12 1 1 17 8 0 0 2 10 1 8 5 1 1 12 4 1 1 17 8 1 1 2 3 0 0 18 5	1						-				-	-			•	0
2 4 0 0 7 10 1 1 12 10 1 1 17 10 1 1 2 2 0 0 7 2 1 1 12 2 1 1 17 2 1 1 2 1 1 7 2 1 1 12 2 1 1 17 2 1 1 2 5 1 1 7 8 1 1 12 9 0 1 17 9 1 1 2 9 0 0 7 9 1 1 12 3 1 1 17 3 1 1 2 6 0 0 7 7 1 1 12 5 1 1 17 3 1 1 2 6 0 0 7 7 1 1 12 5 1 1 1 1 3	1			•				-			-					0
2 7 1 1 7 6 0 0 12 2 1 1 17 5 1 1 2 1 1 7 4 0 0 12 8 1 1 17 9 1 1 2 5 1 1 7 4 0 0 12 8 1 17 9 1 1 2 8 1 1 7 1 1 12 3 1 17 7 1 1 2 9 0 0 7 3 1 1 12 3 1 1 17 3 1 1 2 3 0 0 7 5 1 1 12 4 1 1 17 4 1 1 3 4 0 0 8 1 1 13 4 0 0 18 1 1 1 3 1 1 <t< td=""><td></td><td></td><td></td><td>•</td><td></td><td></td><td></td><td></td><td></td><td></td><td></td><td>•</td><td></td><td></td><td>•</td><td>0</td></t<>				•								•			•	0
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2 1				•							-	•				1
2 5 1 1 7 8 1 1 12 9 0 1 17 9 1 1 2 9 0 0 7 9 1 1 12 7 1 1 17 7 1 1 2 9 0 0 7 9 1 1 12 3 1 1 17 7 1 1 2 6 0 0 7 7 1 1 12 5 1 1 17 8 1 1 3 4 0 0 7 7 1 1 12 5 1 1 17 8 1 1 3 10 1 8 10 0 13 1 1 11 18 1 1 1 3 10 1 8 7 1 1 13 3 0 0 18 1 1 1 11 1				0			-				-	-			•	1
2 8 1 1 7 1 1 12 7 1 1 17 6 0 0 2 10 1 0 7 3 1 1 12 3 1 1 17 3 1 1 2 10 1 0 7 3 1 1 12 1 1 17 3 1 1 2 3 0 0 7 5 1 1 12 4 1 1 17 8 1 1 3 3 0 0 7 5 1 1 12 4 1 1 17 8 1 1 3 3 0 0 13 1 1 13 3 0 18 1 1 1 3 5 0 0 8 1 1 13 3 0 0 18 2 0 1 3 5 0			1	1			0	0			-	1			•	1
2 9 0 0 7 9 1 1 12 3 1 1 17 7 1 1 2 6 0 0 7 7 1 1 12 5 1 1 17 4 1 1 2 3 0 0 7 5 1 1 12 4 1 1 17 8 1 1 3 4 0 0 8 10 0 13 6 0 0 18 1 1 1 3 8 0 1 8 8 0 0 13 6 0 0 18 1 1 1 3 6 1 1 8 3 0 0 13 3 0 0 18 1 </td <td></td> <td></td> <td>1</td> <td>1</td> <td></td> <td></td> <td>1</td> <td>1</td> <td></td> <td></td> <td>0</td> <td>1</td> <td></td> <td></td> <td>•</td> <td>1</td>			1	1			1	1			0	1			•	1
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2 6 0 0 7 7 1 1 12 5 1 1 17 4 1 1 3 4 0 0 8 10 0 0 13 1 1 0 18 9 1 1 3 8 0 1 8 5 0 0 13 2 0 0 18 1 1 1 3 10 1 1 8 5 0 0 13 2 0 0 18 1							-	1				1			•	1
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$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$	4	9	1	1	9	9	0	0	14	5	1	1	19	10	0	1
5 10 1 1 10 2 1 1 15 1 1 0 20 6 0 0 5 8 1 0 10 10 1 15 3 1 1 20 5 0 0 5 9 1 1 10 5 0 0 15 8 0 0 20 2 0 0 5 9 1 1 10 5 0 0 15 8 0 0 20 2 0 0 5 3 0 0 10 9 0 0 15 7 1 1 20 9 1 0 5 6 1 1 10 8 1 1 15 10 1 1 20 7 0 0 5 5 1 1 10 7 0 0 15 2 1 1 20 1 1 0			1	1			1	1		5	1	1				
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	5	7	0	0	10	3	1	1	15	9	0	1	20	10	0	0

D – In Table A1 we show results regarding average earnings, considering that under the two treatments there was a different realization of correct signal, namely, only the probability to get a correct signal was the same (p = 0.75) but it may be that differences in theoretical and experimental earnings were due to chance rather than systematic difference in behaviour. In order to compare average earnings taking into account this caveat, we standardise our results dividing figures in Table 3 by the frequency of correct signals, position by position:

	Theoretic	al earnings		Experimen	tal earnings	
Position in the queue	т1	т2	Т	1	Т	2
			S1	S2	S1	S2
1	0.01875	0.02105	0.01875	0.02222	0.02237	0.02318
2	0.01912	0.02333	0.01618	0.02206	0.02167	0.02462
3	0.02083	0.03393	0.02167	0.01875	0.03036	0.02318
4	0.01949	0.01974	0.01618	0.01875	0.01316	0.02462
5	0.02148	0.02667	0.01719	0.02667	0.02833	0.02667
6	0.02865	0.02578	0.02083	0.02167	0.02344	0.025
7	0.02292	0.02578	0.02	0.00312	0.02344	0.02647
8	0.01910	0.02539	0.01667	0.02692	0.025	0.03077
9	0.02865	0.02361	0.03125	0.025	0.02083	0.02812
10	0.02148	0.03269	0.02344	0.01964	0.03077	0.03214
TOTAL	0.22046	0.257975	0.20214	0.2048	0.239361	0.264777

Table A1 Standardized average earnings for each position and for each treatment

We can observe substantially the same results. At the individual level, for almost all the position, we can still observe higher average earnings under T2, except for position 6 and position 9 for theoretical earnings and for some positions for experimental earnings. However, if we consider average earnings for each treatment, averaging results from different sessions, under T2 earning are higher throughout. At the aggregate level, we can confirm all previous results: the two treatments are statistically different (Wilcoxon rank-sum test for theoretical earnings: -2.043, p-value = .0410; for experimental earnings: -2.570, p-value = .0102). Still, each experimental treatment could be considered not statistically different from its theoretical counterpart (Wilcoxon rank-sum test for BHW and T1: 0.416, p-value = .6772; for our specification and T2: 0.151, p-value = .8798).

E - The aim of this section is to study whether our sample may be divided into two sub-samples: the first four positions and the later ones. We intend to do so performing a Chow test. Before performing this test, we need to verify if a linear probability model (LPM) provides a good approximation for probit estimates (for further details, Wooldridge, 2005). We compare estimates separately from the two models and from each sub-sample in Table A2.

					Cho	ow Test		
Dep. Variable: Correct	Probit (ent	tire)	LPM (entire)	LPM (p	os. 1-4)	LPM (p	os. 5-10)
	Coeff.	<i>p</i> -value	Coeff.	<i>p</i> -value	Coeff.	<i>p</i> -value	Coeff.	<i>p</i> -value
Time	.06416	0.073	.01980	0.063	.01247	0.483	.02526	0.056
Time ²	00236	0.158	00073	0.137	00044	0.595	00097	0.114
Correctwon	17292	0.118	05099	0.119	.04087	0.499	07677	0.062
Correctlost	41712	0.026	13133	0.023	13141	0.199	09515	0.186
т2	.25983	0.009	.07609	0.009	.16531	0.007	.01930	0.628
Constant	.39705	0.019	.66201	0.000	.62892	0.000	.69556	0.000
Pseudo R ²	0.0198							
Adjusted R ²			0.0	152	0.0	208	0.0	083
NOBs	800		80	00	32	20	48	80

Table A2 Binary choice models: probit and LPM comparison

It is straightforward to note that LPM provides consistent results with probit estimates (signs of the coefficient are the same across models and the same variables are statistically significant), even if different in magnitude (due to different assumptions on the error terms).

At this point, we can consider the underlying model as linear and, hence, proceed to perform the Chow test. Under the null hypothesis, there should be no significant difference between the two subsamples. We use the F statistic to test this hypothesis.

We report in table A3 the residual sums of squares for the separate regressions (RSS1 and RSS2) and the residual sum of squares of the pooled sample regression. We compare the value of the total residual sum of squares (obtained by summing RSS1 and RSS2) with the residual sum of squares from the pooled sample regression.

Residual Sum of Squares			
Regression	Positions 1-4 (N.	Positions 5-10	Total
	320)	(N. 480)	(N. 800)
	RSS1	RSS ₂	RSS1+ RSS2
Separate	59.24791	74.18420	133.43211
Pooled			134.88850
F-test	1.433		

Table A3 The Chow test

The F-statistic is 1.433 and the critical value of F(6, 788) at 5% of significance is 2.10. Hence, we conclude that the pooled regression model is an adequate specification. Consequently, there is no statistical difference across the two sub-samples.

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